

## Continuous-Time Markov Chains (2)

- **Alternate definition and Properties of a Poisson Process**
- A continuous-time stochastic process  $\{N(t), t \geq 0\}$  counting the number of events (e.g. arrivals) by time  $t$  is said to be a Poisson process with rate  $\lambda > 0$ , if
  - (i)  $N(0) = 0$ .
  - (ii)  $N(t)$  has *independent increments*. The number of events that occurs in disjoint time intervals are independent.
  - (iii) The number of events that occur in a time interval of length  $t$  is Poisson distributed with mean  $\lambda t$ .
- Obviously, since  $N(t)$  counts the number of events, then  $N(t)$  takes on nonnegative integer values and  $N(t) \geq N(s)$  for  $t > s$ .
- Because a Poisson process is also a pure birth process with birth rates all equal to  $\lambda$ , *the inter-event time is exponentially distributed with mean  $1/\lambda$* .
- Suppose each event in a Poisson process,  $N(t)$ , with rate  $\lambda$  can be classified into type I w.p.  $p$  and type II w.p.  $1-p$ .
- Then, the number of type I and type II,  $N_1(t)$  and  $N_2(t)$  are independent Poisson processes with rates  $\lambda p$  and  $\lambda(1-p)$ .
- This is the *decomposition property* of the Poisson process.

- **Example 5**

- Suppose cars arrive to a gas station according to a Poisson process with rate 5 per hour.

- What is the probability that 3 cars arrive in an hour?

- Number of cars in an hour,  $N(1)$ , is Poisson distributed with mean 5. Then,

$$P\{N(1) = 3\} = e^{-5}(5)^3/3! = 0.14 .$$

- What is the probability that 2 cars arrive in 15 minutes?

- Number of cars in an 15 minutes,  $N(1/4)$ , is Poisson distributed with mean  $5/4 =$ . Then,

$$P\{N(1/4) = 2\} = e^{-5/4}(5/4)^2/2! = 0.224.$$

- What is the expected time before that the third car in an hour arrives?

- Inter-arrival times are exponential with mean  $1/5$  hours. Then, with the memoryless property, the expected time till third is  $3/5$  hours = 36 minutes.

- What is the probability that the station, starting empty, has no cars for 30 minutes after opening?

- Let  $A$  be the inter-arrival time.  $A$  is exponential with mean  $1/5$  hours. The desired probability is

$$P\{A > 1/2\} = e^{-5/2} = 0.082.$$

- **Example 6**

- Customers arrive to a system according to a Poisson process with rate  $\lambda$  and if each customer is a man w.p. 0.5 and a woman w.p. 0.5.
- Characterize the arrival process of men into the system.
- It's a Poisson process with rate  $0.5\lambda$ .

- **Example 7**

- Cars arrive to an intersection according to a Poisson process with rate  $\lambda$ . A policeman blocks one way and directs the cars to the other way. On average, the policeman directs half of the cars to street A and the other half to street B.
- Characterize the arrival process of cars into street A. Is it Poisson?
- It's not a Poisson process.

- **Limiting Probabilities**

- Similar to the discrete cases, we define limiting probabilities for a CTMC as

$$P_j = \lim_{t \rightarrow \infty} P_{ij}(t).$$

- Similar to the discrete case also, these probabilities can be interpreted as the long-run fraction of time spent in state  $j$ .

- The limiting probabilities exist under the following conditions: (i) all states communicate; and (ii) all states are *positive recurrent* meaning that the expected time to return to a state upon leaving it is finite.
- Assuming that the limiting probabilities exist, they can be determined by Kolmogorov's equations (Theorem 1).
- Letting  $t \rightarrow \infty$  in Theorem 1, implies that

$$\lim_{t \rightarrow \infty} \frac{\partial P_{ij}(t)}{\partial t} = \sum_{k \neq j} q_{kj} \lim_{t \rightarrow \infty} P_{ik}(t) - v_j \lim_{t \rightarrow \infty} P_{ij}(t) \Rightarrow 0 = \sum_{k \neq j} q_{kj} P_k - v_j P_j$$

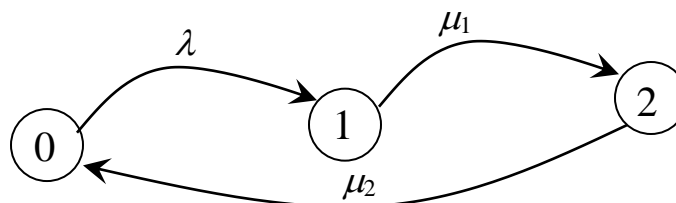
- Therefore,

$$v_j P_j = \sum_{k \neq j} q_{kj} P_k$$

- This equation has an interesting and useful interpretation. The left hand side,  $v_j P_j$ , is the flow out of state  $j$ , and the right hand side,  $\sum_{k \neq j} q_{kj} P_k$ , is the flow into state  $j$ .
- This is a *flow balance equation* (flow out = flow in).

- **Example 8**

- For the shoeshine shop of Example 1, what is the fraction of time that the shop is busy?



- The desired probability is  $P_1 + P_2$  or  $1 - P_0$ .
- The limiting probabilities are given by the following flow balance equations.

State 0.  $\lambda P_0 = \mu_2 P_2$ .

State 1.  $\mu_1 P_1 = \lambda P_0$ .

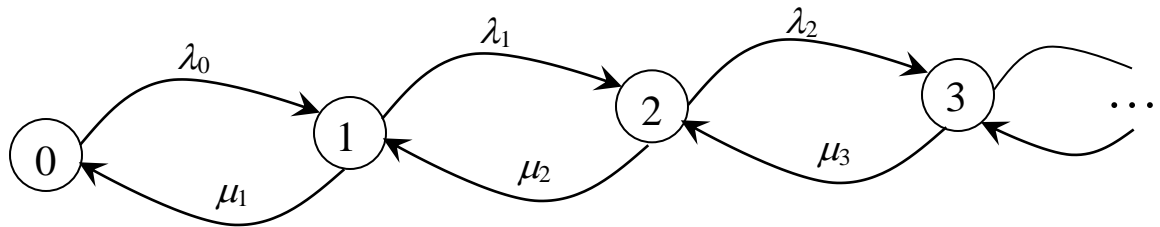
State 2.  $\mu_2 P_2 = \mu_1 P_1$ .

- Therefore,  $P_1 = (\lambda/\mu_1)P_0$ ,  $P_2 = (\lambda/\mu_2)P_0$ . Noting that  $P_0 + P_1 + P_2 = 1$  implies that  $P_0 = 1/[1 + (\lambda/\mu_1) + (\lambda/\mu_2)]$ .

- The desired probability is

$$1 - P_0 = [(\lambda/\mu_1) + (\lambda/\mu_2)] / [1 + (\lambda/\mu_1) + (\lambda/\mu_2)].$$

- **Limiting Probabilities for a Birth-Death Process**



- The flow balance equations are

State 0.  $\lambda_0 P_0 = \mu_1 P_1$ .

State 1.  $\mu_1 P_1 + \lambda_1 P_1 = \lambda_0 P_0 + \mu_2 P_2$ .

State 2.  $\mu_2 P_2 + \lambda_2 P_2 = \lambda_1 P_1 + \mu_3 P_3$ .

State  $n > 2$ .  $\mu_n P_n + \lambda_n P_n = \lambda_{n-1} P_{n-1} + \mu_{n+1} P_{n+1}$ .

- Replacing the first equation in the second gives  $\lambda_1 P_1 = \mu_2 P_2$ .
- Replacing this in the third equation gives  $\lambda_2 P_2 = \mu_3 P_3$ .

➤ Continuing in this manner we find that

$$\lambda_n P_n = \mu_{n+1} P_{n+1}, \quad n \geq 0$$

➤ Then,

$$P_1 = \frac{\lambda_0}{\mu_1} P_0, \quad P_2 = \frac{\lambda_1}{\mu_2} P_1 = \frac{\lambda_1 \lambda_0}{\mu_2 \mu_1} P_0, \quad P_3 = \frac{\lambda_2}{\mu_3} P_2 = \frac{\lambda_2 \lambda_1 \lambda_0}{\mu_3 \mu_2 \mu_1} P_0.$$

➤ In general,

$$P_n = \frac{\lambda_{n-1} \lambda_{n-2} \cdots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \cdots \mu_2 \mu_1} P_0.$$

➤ Since  $\sum_{n=0}^{\infty} P_n = 1$ , it follows that

$$P_0 = \left( 1 + \sum_{n=1}^{\infty} \frac{\lambda_{n-1} \lambda_{n-2} \cdots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \cdots \mu_2 \mu_1} \right)^{-1},$$

$$P_n = \frac{\lambda_{n-1} \lambda_{n-2} \cdots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \cdots \mu_2 \mu_1} \left( 1 + \sum_{n=1}^{\infty} \frac{\lambda_{n-1} \lambda_{n-2} \cdots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \cdots \mu_2 \mu_1} \right)^{-1}, \quad n = 1, 2, \dots$$

➤ A necessary condition for the existence of  $P_n$ s is

$$\sum_{n=1}^{\infty} \frac{\lambda_{n-1} \lambda_{n-2} \cdots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \cdots \mu_2 \mu_1} < \infty.$$